

Introduction to Probability and Statistics - 18.05

Problem set 7 solutions

1. Since the dataset is a realization of a random sample from a Geometric distribution with $p = 1/M$, the likelihood function is

$$\begin{aligned}L(M) &= P(Y_1 = y_1, Y_2 = y_2, \dots, Y_n = y_n) \\&= \left(1 - \frac{1}{M}\right)^{(y_1-1)} \frac{1}{M} \left(1 - \frac{1}{M}\right)^{(y_2-1)} \frac{1}{M} \dots \left(1 - \frac{1}{M}\right)^{(y_n-1)} \frac{1}{M} \\&= \left(1 - \frac{1}{M}\right)^{(-n + \sum_{i=1}^n y_i)} \left(\frac{1}{M}\right)^n \\l(M) &= -n \ln(M) + (-n + \sum_{i=1}^n y_i) \ln\left(1 - \frac{1}{M}\right) \\ \frac{d}{dM} l(M) &= \frac{-n}{M} + (-n + \sum_{i=1}^n y_i) \frac{1}{M(M-1)}\end{aligned}$$

Which equals zero if and only if $M = \bar{y}_n$.

2. Because this is a uniform distribution, $f(x) = \frac{1}{\beta-\alpha}$. So the likelihood function is: $L(\beta) = f(x_1)f(x_2)\dots f(x_n) = \frac{1}{(\beta-\alpha)^n}$ if $\beta \geq \max(x_1, x_2, \dots, x_n)$ and zero otherwise. Therefore, the MLE for β is $\max(x_1, x_2, \dots, x_n)$, because beyond the maximum, the larger the β the smaller $L(\beta)$. I.e. it attains the maximum when $\beta = \max(x_1, x_2, \dots, x_n)$. By a similar argument MLE for α is $\min(x_1, x_2, \dots, x_n)$.
3. (a) Because we have a uniform distribution, we know that $E[Z_i] = 0$ and $Var[Z_i] = \frac{\delta^2}{3}$, thus $E[Z_i^2] = Var[Z_i] + (E[Z_i])^2 = \frac{\delta^2}{3}$. Thus by linearity of expectation $E[X] = \frac{3}{n}(\frac{\delta^2}{3} + \dots + \frac{\delta^2}{3}) = \frac{3}{n} \times n \times \frac{\delta^2}{3} = \delta^2$. So this is a unbiased estimator for δ^2
- (b) Let $g(x) = -\sqrt{x}$, it can be checked by taking double derivative that this is a strictly convex function. Therefore we use Jensen's

inequality

$$\begin{aligned}g(E[X]) &< E[g(x)] \\ -\sqrt{E[X]} &< E[-\sqrt{x}] \\ -\sqrt{\delta} &< -E[\sqrt{x}] \\ -\delta &< -E[\sqrt{x}] \\ \Rightarrow E[\sqrt{x}] &< \delta\end{aligned}$$

Thus it is negatively biased.

4. (a) $E[Y] = E[s\bar{Z}_n + (1-s)\bar{W}_m] = sE[\bar{Z}_n] + (1-s)E[\bar{W}_m]$
 $= s\mu + (1-s)\mu = \mu$ Therefore it is an unbiased estimator.

(b) $MSE(Y) = Var(Y) = s^2Var(\bar{Z}_n) + (1-s)^2Var(\bar{W}_m)$
 $= s^2\frac{\sigma^2}{n} + (1-s)^2\frac{\sigma^2}{m}$

Differentiate this with respect to s , and set it to zero. We get

$$\begin{aligned}\frac{2s}{n} - \frac{2(1-s)}{m} &= 0 \\ \Rightarrow s &= \frac{n}{n+m}\end{aligned}$$

5. (a) To get an unbiased estimate, let $E[T_c] = c(E[X_1] + \dots + E[X_n]) = \frac{1}{\lambda}$

$$\Rightarrow c \times \frac{n}{\lambda} = \frac{1}{\lambda}$$

$$\Rightarrow c = \frac{1}{n}$$

(b)

$$\begin{aligned}MSE &= Var(T_c) + (E(T_c) - \theta)^2 \\ &= c^2n \times \frac{1}{\lambda^2} + (c\frac{n}{\lambda} - \frac{1}{\lambda})^2\end{aligned}$$

Differentiate with respect to c and set it to zero, we get $c = \frac{1}{1+n}$